## FACTOR ANALYSIS

Factor analysis examines the interrelationships among a large number of variables and, then, attempts to explain them in terms of their common underlying dimension

- Common underlying dimensions are referred to as factors
- Data Summarization
- Identify latent dimensions within data set
- Identification and understanding of these underlying dimensions is the goal
- Data Reduction
- Discover underlying dimensions to reduce data to fewer variables so all dimensions are represented in subsequent analyses
- Surrogate variables
- Aggregated scales
- Factor Scores
- Avoid multicollinearity problems
- Improve reliability of aggregated scales

1. Variables must be interrelated 20 unrelated variables=20 factors
2. Sample size Min 50, prefer 100
ii. Min 5 observations/item, prefer 10 observations/item

## Exploratory Factor Analysis (EFA)

- Used to discover underlying structure

Confirmatory Factor Analysis (CFA)

- Used to test whether data fit a priori expectations for data structure
- Structural equations modeling

EFA is a data reduction technique Scientific parsimony
Which items are virtually the same thing
Objective: simplification of items into subset of concepts or measures
Part of construct validation (what are underlying patterns in data?)
EFA assesses dimensionality or homogeneity

Factor (or common factors) analysis
In SPSS known as principal axis factoring
Explain relationship between observed vars in terms of latent vars or factors
Factor is a hypothesized construct
Assumes error in items
Precise math not possible, solved by iteration
Communalities (shared var) on diagonal

Factor - Linear composite. A way of turning multiple measures into one thing.
Factor Score - Measure of one person's score on a given factor.
Factor Loadings - Correlation of a factor score with an item. Variables with high loadings are the distinguishing features of the factor.

Communality - ( $h^{2}$ ) - Variance in a given item accounted for by all factors. Sum of squared factor loadings in a row from factor analysis results. These are presented in the diagonal in common factor analysis

## Primary methods:

Scree rule
Kaiser criterion (eigenvalues > 1)

## Scree Plot

Look for bend in plot
Include factor located right at bend point
Kaiser (or Latent Root) criterion
Eigenvalues greater than 1
Also, 1 is the amount of variance accounted for by a single item ( $r^{2}=1.00$ ). If eigenvalue $<1.00$ then factor accounts for less variance than a single item.

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Example
R matrix (correlation matrix)
BIPr LSat Chol LStr BdWt JSat JStr
```


## BIPr 1.00

```
\begin{tabular}{llllllll} 
LSat & -.18 & 1.00 & & & & & \\
Chol & .65 & -.17 & 1.00 & & & & \\
LStr & .15 & \(\underline{-.45}\) & .22 & 1.00 & & & \\
BdWt & \(\underline{.45}\) & -.11 &. .52 & .16 & 1.00 & & \\
JSat -.21 & \(\underline{.85}\) & -.12 & \(\underline{-.35}-.05\) & & 1.00 & & \\
JStr & .19 & -.21 & .02 & \(\underline{.79}\) & .19 & \(\underline{-.35}\) & 1.00
\end{tabular}
Principal Components Analysis (PCA) Initial Statistics:
Variable Communality * Factor Eigenval \%Var Cum\%
\begin{tabular}{llllll} 
BLPR & 1.00000 * & 1 & 2.85034 & 40.7 & 40.7
\end{tabular}
LSAT 1.00000 * 2 1.74438 24.9 65.6
CHOL 1.00000 * 31.16388 16.6 82.3
LSTR 1.00000 * 4
BDWT 1.00000 * 5 . 44201 6.3 96.6
JSAT 1.00000 * 6 . 20235 2.9 99.5
JSTR 1.00000 * 7 . 03607 . 5 100.0
```


## Example

| Variable | Communality * Factor |  | Eigenval \%Var Cum\% |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| BLPR | 1.00000 | * 1 | 2.85034 | 40.7 | 40.7 |
| LSAT | 1.00000 | 2 | 1.74438 | 24.9 | 65.6 |
| CHOL | 1.00000 | 3 | 1.16388 | 16.6 | 82.3 |
| LSTR | 1.00000 | 4 | . 56098 | 8.0 | 90.3 |
| BDWT | 1.00000 |  | . 44201 | 6.3 | 96.6 |
| JSAT | 1.00000 | 6 | . 20235 | 2.9 | 99.5 |
| JSTR | 1.00000 | 7 | . 03607 | 5 | 100.0 |

Factor Matrix (Unrotated):

|  | Factor 1 | Factor 2 | Factor 3 $\ldots$.. Fac7 |  |
| :--- | ---: | ---: | ---: | :--- |
| LSTR | .73738 | -.32677 | .47575 |  |
| LSAT | -.71287 | .38426 | .52039 |  |
| JSAT | -.70452 | .42559 | .48553 |  |
| JSTR | .64541 | -.32867 | .62912 |  |
| CHOL | .54945 | .68694 | -.10453 |  |
| BDWT | .48867 | .60471 | .13043 |  |
| BLPR | .58722 | .60269 | -.08534 |  |

Eigenvalue $2.850343 \quad 1.74438 \quad 1.16388$

## Example



Eigenvalue 2.08831 .88091 .7893

Factor Scree Plot


Scree comes from a word for loose rock and debris at the base of a cliff!

## Information from EFA

| Msr | F 1 | F 2 | F 3 | $\mathrm{~h}^{2}$ |
| :--- | :---: | :---: | :---: | :---: |
|  | .60 | -.06 | .02 | .36 |
| b | .81 | .12 | -.03 | .67 |
| c | .77 | .03 | .08 | .60 |
| d | .01 | .65 | -.04 | .42 |
| e | .03 | .80 | .07 | .65 |
| f | .12 | .67 | -.05 | .47 |
| g | .19 | -.02 | .68 | .50 |
| h | .08 | -.10 | .53 | .30 |
| i | .26 | -.13 | .47 | .31 |
| Sum Sq Ldng | 1.76 | 1.56 | .98 | Total |
| \% Variance | .195 | .173 | .109 | $47.7 \%$ |
|  | $(1.76 / 9)$ | $(1.56 / 9)$ | $(.98 / 9)$ |  |

A factor loading is the correlation between a factor and an item When factors are orthogonal, factor loadings squared are the amount of variance in one variable explained by that factor (F1 explains 36\% of the variance in Msr a; F3 explains 46\% of the variance in Msr g)

## Information from EFA

| Msr | F 1 | F 2 | F 3 | $\mathrm{~h}^{2}$ |
| :--- | :---: | :---: | :---: | :---: |
| a | .60 | -.06 | .02 | .36 |
| b | .81 | .12 | -.03 | .67 |
| $\ldots$ | $\ldots$ | $\ldots$ | $\ldots$ | $\ldots$ |
| u | .26 | -.13 | .47 | .31 |
| Sum Sq Ldng | 1.76 | 1.56 | .98 | Total |
| \% Variance | .195 | .173 | .09 | $47.7 \%$ |
|  | $(1.76 / 9)$ | $(1.56 / 9)$ | $(.98 / 9)$ |  |

Eigenvalue: Sum of squared loadings down a column (associated with a factor). Total variance in all vars explained by one factor. Factors with eigenvalues less than 1 predict less than the variance of 1 item.
Communality (h²): Variance in a given item accounted for by all factors. Sum of squared loadings across rows. Will equal 1 if you retain all possible factors.

## Information from EFA

FACTOR

| Msr | F1 | F2 | F3 | h $^{2}$ |
| :--- | :---: | :---: | :---: | :---: |
| a | .60 | -.06 | .02 | .36 |
| b | .81 | .12 | -.03 | .67 |
| $\ldots$ | $\ldots$ | $\ldots$ | $\ldots$ | $\ldots$ |
| i | .26 | -.13 | .47 | .31 |


| Sum Sq Ldng | 1.76 | 1.56 | .98 | Total |
| :--- | ---: | :---: | :---: | :--- |
| \% Variance | .195 | .173 | .109 | $47.7 \%$ |
|  | $(1.76 / 9)$ | $(1.56 / 9)$ | $(.98 / 9)$ |  |

Average of all communalities $\left(h^{2} / k\right)=$ proportion of variance in all variables explained by all factors.
If all variables reproduced perfectly by the factors, correlation between original variables equals sum of the products of factor loadings. When not perfect, gives an estimate of the correlation.
e.g. $r_{\text {ab }} \cong\left(.60^{*} .81\right)+\left(-.06^{*} .12\right)+\left(.02^{*}-.03\right) \cong .48$

## Information from EFA

| Msr | F1 | F2 | F3 | $\mathbf{h}^{2}$ |
| :--- | :---: | :---: | :---: | :---: |
| a | .60 | -.06 | .02 | .36 |
| b | .81 | .12 | -.03 | .67 |
| $\ldots$ | $\ldots$ | $\ldots$ | $\ldots$ | $\ldots$ |
| m | . .26 | -.13 | .47 | .31 |
| Sum Sq Ldng | 1.76 | 1.56 | .98 | Total |
| \% Variance | .195 | .173 | .109 | $47.7 \%$ |
|  | $(1.76 / 9)$ | $(1.56 / 9)$ | $(.98 / 9)$ |  |

$1-h^{2}$ is the uniqueness $\rightarrow$ variance of an item not shared with other items. Unique variance could be random error or systematic.
The factor matrix above is after rotation. Eigenvalues computed on the unrotated and unreduced factor loading matrix because we are interested in total variance accounted for in the data. Use of eigenvalues and \% variance accounted for in SPSS not reordered after rotation.

## Important Properties of

- Each factor in turn maximizes variance explained from an R matrix
o For any number of factors obtained, PCs maximize variance explained
o Amount of variance explained by each PC equals the corresponding characteristic root (eigenvalue)
o All characteristic roots of PCs are positive
o Number of PCs derived equal the number of factors need to explain all the variance in $\mathbf{R}$
o The sum of characteristic roots equals the sum of diagonal R elements


## Rotations

- All original PC and PF solutions are orthogonal.
o Once you obtain minimal number of factors, you have to interpret them
o Interpreting original solutions is difficult. Rotation aids interpretation.
- You are looking for simple structure
- Component loadings should be very high for a few vars and near 0 for remaining variables
- Each variable should load highly on only 1 component Unrotated Matrix Rotated Matrix

| VarF1 | F2 | F1 F2 |
| :---: | :---: | :---: |
| a | .75 .63 | .14 .95 |
| $b$ | .69 .57 | .14 .90 |
| $c$ | .80 .49 | .18 .92 |
| $d$ | $85-47$ | 04.0 |

## Rotation

- After rotation, variance accounted for by a factor is spread out. First factor no longer accounts for max variance possible; others get more variance. Total variance accounted for is the same.
- Two types of rotation
- Orthogonal (factors uncorrelated) - Oblique (factors correlated)


## Rotation

o Orthogonal rotation (rigid, 90 degrees) - PCs or PFs remain uncorrelated after transformation

- Varimax - Simplifying column weights to 1s and 0s. Factor has items loading highly, others don't load. Not appropriate if you expect a single factor.
- Quartimax - Simplify to 1 s and 0 s in a row. Item loads high on 1 factor, almost 0 on others. Appropriate if you expect single general factor.
- Equimax. Compromise of Varimax and Quartimax rotations.
- In practice, choice of rotation makes little difference


## Rotation

o Oblique or correlated components (less or more than 90 degrees) - Account for same \% var, but factors correlated

- Some say not meaningful with PCA
- Many factors are theoretically related, so rotation method should not "force" orthogonality
- Allows the loadings to more closely match simple structure
- Correlated solutions will get you closer to simple structure
- Oblimin (Kaiser) and promax are good
- Provides a structure matrix of loadings and a pattern matrix of partial weights - which to interpret?


## Orthogonal Rotation



## Simple Structure (Thurstone)

(1) Each row of factor matrix should have at least one 0 loading
(2) The number of items with 0 loadings equals the number of factors; each column has 1 or more 0 loadings
(3) Items with high loadings on one factor or the other
(4) If there are more than 4 factors, a large portion of items should have zero loadings
(5) For every pair of columns, there should be few cross-loadings
(6) Few if any negative loadings

## Simple Structure

Factor

| Msr | 1 | 2 | 3 |
| :--- | :--- | :--- | :--- |
| a | x | 0 | 0 |
| b | x | 0 | 0 |
| c | X | 0 | 0 |
| d | 0 | x | 0 |
| e | 0 | x | 0 |
| f | 0 | x | 0 |
| g | 0 | 0 | x |
| h | 0 | 0 | x |
| i | 0 | 0 | x |
| j | 0 | 0 | x |

## Oblique Rotation

- Example:

Unrotated Matrix Rotated Matrix


## Orthogonal or Oblique Rotation?

- Nunnally suggests using orthogonal as opposed to oblique rotations
- Orthogonal is simpler
- Leads to same conclusions
- Oblique can be misleading
- Ford et al. suggest using oblique unless orthogonality assumption is tenable


## Interpretation

o Factors usually interpreted by observing which variables load highest on each factor

- a priori criteria for loadings (min .3+)
o Name factor. Always provide factor loading matrix in study.
o Cross-loadings are problematic
- a priori criteria for "large" cross-loading
- decide a priori what you will do
o Factor loadings or summated scales used to define new scale. Can go back to correlation matrix and do not only use factor loadings. Loadings can be inflated.


## PCA and FA

- PCA - No constructs of theoretical meaning assumed; Simple mechanical linear combination. (1s in the diagonal of R)
o FA - assumes underlying latent constructs. Allows for measurement error (communalities in diagonal of R)
- Also PAF or common factors analysis
- PCA uses all the variance. FA uses ONLY shared variance.
- In FA you can have indeterminant (unsolvable) solutions. Have to iterate (computer makes best "guess") to get the solutions.


## FA

- Also known as principal axis factoring or common factor analysis
- Steps
- Estimate communalities of the variables (shared variance)
- Substitute communalities in place of 1 s on diagonal of $\mathbf{R}$
- Perform a principal component analysis on the reduced matrix
- Iterated FA
- Estimate h ${ }^{2}$
- Solve for factor model
- Calculate new communalities
- Substitute new estimates of $h^{2}$ into matrix and redo
- Iterate until communalities don't change much
- Rotate for interpretation


## Estimating Communalities

- Highest correlation of given variable with other variables in data set
- Squared multiple correlations (SMCs) of each variable predicted by all other variables in the data set
- Reliability of the variable
- Because you are estimating and the factors are no longer combinations of actual variables, can get funny results:
- Communalities > 1.00
- Negative eigenvalues
- Negative uniqueness


## Example FA

R matrix (correlation matrix with h²) BIPr LSat Chol LStr BdWt JSat JStr
BIPr . 54

| LSat | -.18 | .89 |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Chol | $\underline{.65}$ | -.17 | .67 |  |  |  |  |
| LStr . 15 | $\frac{-.45}{}$ | .22 | .87 |  |  |  |  |
| BdWt | $\underline{.45}$ | -.11 | .52 | .16 | .41 |  |  |
| JSat | -.21 | $\underline{.85}$ | -.12 | -.35 | -.05 | .86 |  |
| JStr | .19 | -.21 | .02 | $\underline{.79}$ | .19 | $\underline{-.35}$ | .87 |


| Principal Axis Factoring (PAF) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Initial Statistics: |  |  |  |  |  |
| Variable Communality * Factor |  |  | Eigenvalue | \%Var Cum\% |  |
| BLPR | . 53859 | 1 | 2.85034 | 40.7 | 40.7 |
| LSAT | . 88573 * | 2 | 1.74438 | 24.9 | 65.6 |
| CHOL | . 66685 | 3 | 1.16388 | 16.6 | 82.3 |
| LSTR | . 87187 |  | . 56098 | 8.0 | 90.3 |
| BDWT | . 41804 | 5 | . 44201 | 6.3 | 96.6 |
| JSAT | . 86448 * | 6 | . 20235 | 2.9 | 99.5 |
| JSTR | . 86966 | 7 | . 03607 | . 5 | 100.0 |

## FA



Factor Matrix (Unrotated):

|  | Factor 1 | Factor 2 | Factor 3 |
| :--- | ---: | ---: | ---: |
| LSAT | -.75885 | .31104 | .54455 |
| LSTR | .70084 | -.20961 | .36388 |
| JSAT | -.70038 | .31502 | .39982 |
| JSTR | .68459 | -.29044 | .66213 |
| CHOL | .48158 | .74399 | -.07267 |
| BLPR | .48010 | .56066 | -.02253 |
| BDWT | .36699 | .47668 | .08381 |

\%Var Cum\%

## FA

## Principal Axis Factoring (PAF)

Final Statistics:

| Variable | Communality * | Factor Eigenvalue | \%Var | Cum\% |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
| BLPR | $.545355^{*}$ | 1 | 2.62331 | 37.5 | 37.5 |
| LSAT | .96913 * | 2 | 1.41936 | 20.3 | 57.8 |
| CHOL | $.79071^{*}$ | 3 | 1.04004 | 14.9 | 72.6 |
| LSTR | $.66752^{*}$ |  |  |  |  |
| BDWT | $.36893^{*}$ |  |  |  |  |
| JSAT | .74962 * |  |  |  |  |
| JSTR | .99144 * |  |  |  |  |

Rotated Factor Matrix (VARIMAX):
Factor 1 Factor 2 Factor 3

| LSAT | .96846 | -.10483 | -.14223 |
| :--- | :---: | :---: | :---: |
| JSAT | .83532 | -.07092 | -.21643 |
| CHOL | -.08425 | .88520 | -.00547 |
| BLPR | -.11739 | .72364 | .08898 |
| BDWT | -.00430 | .59379 | .12778 |
| JSTR | -.10474 | .07011 | .98770 |
| LSTR | -.28514 | .15273 | .75026 |

## Logic of FA

 BIPr LSat Chol LStr BdWt JSat JStrHow many? What are the factors?


BIPr LSat Chol LStr BdWt JSat JStr

## PCA vs. FA

o Pros \& Cons:

- Pro PCA: has solvable equations. "Math is right".
- Con PCA: Lumping garbage together. Also, no underlying concepts.
- Pro FA: considers role of measurement error, gets at concepts.
- Con FA: doing mathematical gymnastics.
o Practically: Usually not much difference
- PCA will tend to converge more consistently
- FA is more meaningful conceptually

PCA vs. FA
o Situations where you might want to use FA:

- Where there are 12 or fewer variables (diagonal will have a large impact)
- Where the correlations between the variables are small, then diagonals will have a large impact
o If you have clear factor structure, won't make much difference
o Otherwise:
- PCA will tend to overfactor
- If doing exploratory analysis, may not mind overfactoring


## Using FA Results

o Single surrogate measure - choose a single item with a high loading to represent factor

- Summated Scale*
- Form a composite from items loading on same factor
- Average all items that load on a factor (unit weighting)
- Calculate the alpha for the reliability
- Name the scale/construct
- Factor Scores
- Composite measures for each factor were computed for each subject
- Based on all factor loadings for all items
- Not easily replicated


## Reporting

- If you create a factor based scale, describe the process
o Factor analytic study, report:
- Theoretical rationale for EFA
- Detailed description of subjects and items, including descriptive stats
- Correlation matrix
- Methods used (PCA/FA, communality estimates, factor extraction, rotation)
- Criteria employed for number of factors and meaningful loadings
- Factor matrix (aka pattern matrix)


## Confirmatory Factor Analysis

o Part of construct validation process (do the data conform to expectations regarding the underlying patterns?)

- Use SEM packages to perform CFA
- EFA with specified number of factors for a criterion is NOT a CFA
- Basically start with a correlation matrix and expected relationships
o Look at whether expected relationships can reproduce the correlation matrix well
- Tested with chi-square goodness of fit. If significant, data don't fit expected structure. No confirmation.
o Alternative measures of fit available.


## Logic of CFA

Let's say I believe:


BIPr LSat Chol LStr BdWt JSat JStr But the reality is
Phys Hlth


BIPr LSat Chol LStr BdWt JSat JStr Data won't confirm expected structure

## Example

R matrix (correlation matrix)
BIPr LSat Chol LStr BdWt JSat JStr
BIPr 1.00
LSat -. 181.00
$\begin{array}{lllll}\text { Chol } & \underline{65} & -.17 & 1.00 \\ \text { LStr } & .15 & -.45 & .22 & 1.00\end{array}$
BdWt . 45 -. $11 \quad .52$. $16 \quad 1.00$
JSat $\quad-.21 \quad . \quad .85-.12-. .35-.05 \quad 1.00$


BIPr LSat Chol LStr BdWt JSat JStr

