Simple Linear Regression Analysis

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The Simple Linear Regression Model and the Least Squares Point Estimates

- The **dependent** (or response) variable is the variable we wish to understand or predict
- The **independent** (or predictor) variable is the variable we will use to understand or predict the dependent variable
- **Regression analysis** is a statistical technique that uses observed data to relate the dependent variable to one or more independent variables
- The objective is to build a regression model that can describe, predict and control the dependent variable based on the independent variable

Form of The Simple Linear Regression Model

- $y = \beta_0 + \beta_1 x + \epsilon$
- $\mu_y = \beta_0 + \beta_1 x + \varepsilon$ is the mean value of the dependent variable y when the value of the independent variable is x
- β_0 is the y-intercept; the mean of y when x is 0
- β₁ is the slope; the change in the mean of y per unit change in x
- ε is an error term that describes the effect on y of all factors other than x



Regression Terms

- β_0 and β_1 are called regression parameters
- β_0 is the y-intercept and β_1 is the slope
- We do not know the true values of these parameters
- So, we must use sample data to estimate them
- b_0 is the estimate of β_0 and b_1 is the estimate of β_1

The Simple Linear Regression Model Illustrated

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The Least Squares Point Estimates

- Estimation/prediction equation $\hat{y} = b_0 + b_1 x$
- Least squares point estimate of the slope β_1 $b_1 = \frac{SS_{xy}}{SS_{xx}}$ $SS_{xy} = \sum (x_i - \overline{x})(y_i - \overline{y}) = \sum x_i y_i - \frac{(\sum x_i)(\sum y_i)}{n}$

$$SS_{xx} = \sum (x_i - \bar{x})^2 = \sum x_i^2 - \frac{(\sum x_i)^2}{n}$$

• Least squares point estimate of y-intercept β_0

$$b_0 = \overline{y} - b_1 \overline{x}$$
 $\overline{y} = \frac{\sum y_i}{n}$ $\overline{x} = \frac{\sum x_i}{n}$

The Tasty Sub Shop Case #1

| Уi | X _i | x_i^2 | x _i y _i |
|---------------------|--------------------|--------------------------------|-------------------------------|
| 527.1 | 20.8 | $(20.8)^2 = 432.64$ | (20.8)(527.1) = 10963.68 |
| 548.7 | 27.5 | $(27.5)^2 = 756.25$ | (27.5)(548.7) = 15089.25 |
| 767.2 | 32.3 | (32.3) ² = 1,043.29 | (32.3)(767.2) = 24780.56 |
| 722.9 | 37.2 | (37.2) ² = 1,383.84 | (37.2)(722.9) = 26891.88 |
| 826.3 | 39.6 | (39.6) ² = 1,568.16 | (39.6)(826.3) = 32721.48 |
| 810.5 | 45.1 | (45.1) ² = 2,034.01 | (45.1)(810.5) = 36553.55 |
| 1040.7 | 49.9 | $(49.9)^2 = 2,490.01$ | (49.9)(1040.7) = 51930.93 |
| 1033.6 | 55.4 | (55.4) ² = 3,069.16 | (55.4)(1033.6) = 57261.44 |
| 1090.3 | 61.7 | (61.7) ² = 3,806.89 | (61.7)(1090.3) = 67271.51 |
| 1235.8 | 64.6 | $(64.6)^2 = 4,173.16$ | (64.6)(1235.8) = 79832.68 |
| $\sum y_i = 8603.1$ | $\sum x_i = 434.1$ | $\sum x_i^2 = 20,757.41$ | $\sum x_i y_i = 403,296.96$ |

The Tasty Sub Shop Case #2

- From last slide,
 - $\Sigma y_i = 8,603.1$
 - $\Sigma x_i = 434.1$
 - $\Sigma x_{i}^{2} = 20,757.41$
 - $\Sigma x_i y_i = 403,296.96$
- Once we have these values, we no longer need the raw data
- Calculation of b₀ and b₁ uses these totals

The Tasty Sub Shop Case #3 (Slope b₁)

$$SS_{xy} = \sum x_i y_i - \frac{\left(\sum x_i\right)\left(\sum y_i\right)}{n}$$

= 403,296.96 - $\frac{(434.1)(8,603.1)}{10}$ = 29,836.389
$$SS_{xx} = \sum x_i^2 - \frac{\left(\sum x_i\right)^2}{n}$$

= 120,757.41 - $\frac{(434.1)^2}{10}$ = 1,913.129

$$b_1 = \frac{SS_{xy}}{SS_{xx}} = \frac{29,836.389}{1,913.129} = 15.596$$

The Tasty Sub Shop Case #4 (y-Intercept b_0)

$$\overline{y} = \frac{\sum y_i}{n} = \frac{8,603.1}{10} = 860.31$$
$$\overline{x} = \frac{\sum x_i}{n} = \frac{434.1}{10} = 43.41$$

$$b_0 = \overline{y} - b_1 \overline{x}$$

= 860.31 - (15.596)(43.41)
= 183.31



- Prediction (x = 20.8)
- $\hat{y} = b_0 + b_1 x = 183.31 + (15.59)(20.8)$
- ŷ = 507.69
- Residual is 527.1 507.69 = 19.41

Figure 14.5

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Model Assumptions and the Standard Error

1. Mean of Zero

At any given value of x, the population of potential error term values has a mean equal to zero

2. Constant Variance Assumption

At any given value of x, the population of potential error term values has a variance that does not depend on the value of x

3. Normality Assumption

At any given value of x, the population of potential error term values has a normal distribution

4. Independence Assumption

Any one value of the error term ε is statistically independent of any other value of ε





Sum of Squares

- Sum of squared errors $SSE = \sum e_i^2 = \sum (y_i - \hat{y}_i)^2$
- Mean square error
 - $\circ\,$ Point estimate of the residual variance σ^2

$$s^2 = MSE = \frac{SSE}{n-2}$$

- Standard error
 - \circ Point estimate of residual standard deviation σ

$$s = \sqrt{MSE} = \sqrt{\frac{SSE}{n-2}}$$

Testing the Significance of the Slope and y-Intercept

- A regression model is not likely to be useful unless there is a significant relationship between x and y
- To test significance, we use the null hypothesis:

H0: $\beta_1 = 0$

• Versus the alternative hypothesis:

Ha: $\beta_1 \neq 0$

| Testing the #2 | e Significai | nce of the Slope |
|-----------------------|--------------------------------|--|
| <u>Alternative</u> | <u>Reject H₀ If</u> | <u><i>p</i>-Value</u> |
| $H_a: \beta_1 > 0$ | $t > t_{\alpha}$ | Area under t distribution right of t |
| $H_a: \beta_1 < 0$ | $t < -t_{\alpha}$ | Area under t distribution left of t |
| $H_a: \beta_1 \neq 0$ | $ t > t_{\alpha/2}^{*}$ | Twice area under t distribution right of t |

* That is $t > t_{\alpha/2}$ or $t < -t_{\alpha/2}$

Testing the Significance of the Slope #3

• Test Statistics $t = \frac{b_1}{s_{b_1}}$ where $s_{b_1} = \frac{s}{\sqrt{SS_{xx}}}$

- 100(1- α)% Confidence Interval for β_1 [$b_1 \pm t_{\alpha/2} S_{b1}$]
- t_{α} , $t_{\alpha/2}$ and p-values are based on n–2 degrees of freedom

Confidence and Prediction Intervals

- The point on the regression line corresponding to a particular value of x_0 of the independent variable x is $\hat{y} = b_0 + b_1 x_0$
- It is unlikely that this value will equal the mean value of y when x equals x₀
- Therefore, we need to place bounds on how far the predicted value might be from the actual value
- We can do this by calculating a confidence interval mean for the value of y and a prediction interval for an individual value of y



Distance Value

- Both the confidence interval for the mean value of y and the prediction interval for an individual value of y employ a quantity called the distance value
- The distance value for a particular value x_0 of x is

$$\frac{1}{n} + \frac{(x_0 - \overline{x})^2}{SS_{xx}}$$

- The distance value is a measure of the distance between the value x_0 of x and \bar{x}
- Notice that the further x₀ is from \bar{x} , the larger the distance value

A Confidence and Prediction Interval for a Mean Value of y

- Assume that the regression assumption holds
- The formula for a 100(1-α) confidence interval for the mean value of y is as follows:

 $[\hat{y} \pm t_{\alpha/2} s \sqrt{\text{Distance value}}]$

 The formula for a 100(1-α) prediction interval for an individual value of y is as follows:

 $[\hat{y} \pm t_{\alpha/2} s \sqrt{1 + \text{Distance value}}]$

• This is based on n-2 degrees of freedom



Which to Use?

- The prediction interval is useful if it is important to predict an individual value of the dependent variable
- A confidence interval is useful if it is important to estimate the mean value
- The prediction interval will always be wider than the confidence interval

14.5 Simple Coefficient of Determination and Correlation

- How useful is a particular regression model?
- One measure of usefulness is the simple coefficient of determination
- It is represented by the symbol r²

Calculating The Simple Coefficient of Determination

- 1. Total variation is given by the formula $\Sigma(y_i-\bar{y})^2$
- 2. Explained variation is given by the formula $\Sigma(\hat{y}_i \bar{y})^2$
- 3. Unexplained variation is given by the formula $\Sigma(y_i \hat{y})^2$
- 4. Total variation is the sum of explained and unexplained variation
- 5. r² is the ratio of explained variation to total variation

The Simple Correlation Coefficient

• The simple correlation coefficient measures the strength of the linear relationship between y and x and is denoted by r

$$r = +\sqrt{r^2}$$
 if b_1 is positive, and
 $r = -\sqrt{r^2}$ if b_1 is negative

• Where b₁ is the slope of the least squares line

Different Values of the Correlation Coefficient

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Testing the Significance of the Population Correlation Coefficient

- The simple correlation coefficient (r) measures the linear relationship between the observed values of x and y from the sample
- The population correlation coefficient (p) measures the linear relationship between all possible combinations of observed values of x and y
- r is an estimate of p

Testing ρ

• We can test to see if the correlation is significant using the hypotheses

 $\begin{array}{l} H_0:\,\rho=0\\ H_a:\,\rho\neq 0 \end{array}$

• The statistic is
$$t = \frac{r \cdot \sqrt{n-2}}{\sqrt{1-r^2}}$$

• This test will give the same results as the test for significance on the slope coefficient b₁



An F Test for Model

• For simple regression, this is another way to test the null hypothesis

 $H_0: \beta_1 = 0$

- This is the only test we will use for multiple regression
- The F test tests the significance of the overall regression relationship between x and y



Mechanics of the F Test

- To test $H_0: \beta_1 = 0$ versus Ha: $\beta_1 \neq 0$ at the α level of significance
- Test statistics based on F

 $F = \frac{\text{Explained variation}}{(\text{Unexplain ed variation })/(n - 2)}$



- Reject H_0 if $F(model) > F_\alpha$ or p-value $< \alpha$
- F_{α} is based on 1 numerator and n-2 denominator degrees of freedom

The QHIC Case

- Quality Home Improvement Center (QHIC) operates five stores
- Wish to study relationship between home value and yearly expenditure on home upkeep
- Random sample of 40 homeowners
 - Intercept = -348.3921
 - Slope 7.2583



Residual Analysis

- Checks of regression assumptions are performed by analyzing the regression residuals
- Residuals (e) are defined as the difference between the observed value of y and the predicted value of y, e = y - ŷ

• Note that e is the point estimate of ε

- If regression assumptions valid, the population of potential error terms will be normally distributed with mean zero and variance σ^2
- Different error terms will be statistically independent

Residual Analysis #2

- Residuals should as if they are randomly and independently selected from normal populations with mean zero and variance σ^2
- With any real data, assumptions will not hold exactly
- Mild departures do not affect our ability to make statistical inferences
- In checking assumptions, we are looking for pronounced departures from the assumptions
- So, only require residuals to approximately fit the description above



Residual Plots

- 1. Residuals versus independent variable
- 2. Residuals versus predicted y's
- 3. Residuals in time order (if the response is a time series)

Constant Variance Assumptions

- To check the validity of the constant variance assumption, examine residual plots against
 - The x values
 - The predicted y values
 - Time (when data is time series)
- A pattern that fans out says the variance is increasing rather than staying constant
- A pattern that funnels in says the variance is decreasing rather than staying constant
- A pattern that is evenly spread within a band says the assumption has been met

Constant Variance Visually



Assumption of Correct Functional Form

- If the relationship between x and y is something other than a linear one, the residual plot will often suggest a form more appropriate for the model
- For example, if there is a curved relationship between x and y, a plot of residuals will often show a curved relationship

Normality Assumption

- If the normality assumption holds, a histogram or stem-and-leaf display of residuals should look bell-shaped and symmetric
- Another way to check is a normal plot of residuals
 - Order residuals from smallest to largest
 - Plot $e_{(i)}$ on vertical axis against $z_{(i)}$
 - $Z_{(i)}$ is the point on the horizontal axis under the z curve so the area under this curve to the left is (3i-1)/(3n+1)
- If the normality assumption holds, the plot should have a straight-line appearance

Independence Assumption

- Independence assumption is most likely to be violated when the data are time-series data
 - If the data is not time series, then it can be reordered without affecting the data
 - Changing the order would change the interdependence of the data
- For time-series data, the time-ordered error terms can be autocorrelated
 - Positive autocorrelation is when a positive error term in time period i tends to be followed by another positive value in i+k
 - Negative autocorrelation is when a positive error term in time period i tends to be followed by a negative value in i+k
- Either one will cause a cyclical error term over time

Independence Assumption Visually

